

## Solutions to “A Brief Introduction to Stochastic Calculus”

These are some solutions I have written to exercises from these popular [notes](#) from Columbia University’s *IEOR E4706: Foundations of Financial Engineering*. I personally found the notes very helpful for picking up introductory stochastic calculus (e.g. Brownian motions, stochastic integrals, Itô’s lemma) with minimal measure theory background.

Please email [anish.lakapragada@yale.edu](mailto:anish.lakapragada@yale.edu) for any questions or errors.

### Exercise 1: Conditional expectations as martingales

Let  $Z$  be a random variable and set  $X_t := \mathbb{E}[Z \mid \mathcal{F}_t]$ . Show that  $X_t$  is a martingale.

Leaving aside the first property from Definition 3, we show the second property. Namely  $\forall t, s \geq 0$ , we have:

$$\mathbb{E}[X_{t+s} \mid \mathcal{F}_t] = \mathbb{E}[\mathbb{E}[Z \mid \mathcal{F}_{t+s}] \mid \mathcal{F}_t]$$

but  $\mathcal{F}_t \subset \mathcal{F}_{t+s}$  and so by the tower property  $\mathbb{E}[X_{t+s} \mid \mathcal{F}_t] = \mathbb{E}[Z \mid \mathcal{F}_t] = X_t \implies X_t$  is a martingale.

### Exercise 2: Martingale Property of Stochastic Integrals of an Elementary Process

Check that  $Y_t(\omega) := \int_0^t X_s(\omega) dW_s(\omega)$  is indeed a martingale when  $X_t(\omega)$  is an elementary process.

Before showing that  $Y_t(\omega)$  is a martingale, we first provide our canonical definition of elementary process  $X_t(\omega) := \sum_i e_i(\omega) I_{[t_i, t_{i+1})}(t)$  where  $\{e_k(\omega)\}_{k=0}^n$  and  $\{t_k\}_{k=0}^n$  are defined as in Definition 6.

We now check that  $Y_t(\omega)$  is a martingale, this time with both properties in Definition 3. We first start by showing the more interesting second property. Pick  $t, s \geq 0$ . We apply Definition 7:

$$Y_{t+s}(\omega) = \int_0^{t+s} X_s(\omega) dW_s(\omega) = \sum_{i=0}^{n-1} e_i(\omega) [W_{t_{i+1} \wedge (t+s)}(\omega) - W_{t_i \wedge (t+s)}(\omega)]$$

where<sup>a</sup>  $x \wedge y = \min(x, y)$ . We now split this summation based on index  $j$  where  $t_j \leq t \leq t_{j+1}$ :

$$\begin{aligned} Y_{t+s}(\omega) &= \underbrace{\sum_{i=0}^{j-1} e_i(\omega) [W_{t_{i+1}}(\omega) - W_{t_i}(\omega)]}_{Y_t(\omega)} + \sum_{i=j}^{n-1} e_i(\omega) [W_{t_{i+1} \wedge (t+s)}(\omega) - W_{t_i \wedge (t+s)}(\omega)] \\ \implies \mathbb{E}[Y_{t+s}(\omega) \mid \mathcal{F}_t] &= Y_t(\omega) + \mathbb{E}\left[\sum_{i=j}^{n-1} e_i(\omega) [W_{t_{i+1} \wedge (t+s)}(\omega) - W_{t_i \wedge (t+s)}(\omega)] \mid \mathcal{F}_t\right] \end{aligned}$$

where these expectations are over  $\omega \in \Omega$ . To finish,  $\mathbb{E}[W_{t_{i+1} \wedge (t+s)}(\omega) - W_{t_i \wedge (t+s)}(\omega)] = 0 \implies \mathbb{E}[\sum_{i=j}^{n-1} e_i(\omega)[W_{t_{i+1} \wedge (t+s)}(\omega) - W_{t_i \wedge (t+s)}(\omega)] \mid \mathcal{F}_t] = 0$ , and so we have demonstrated the martingale property  $\mathbb{E}[Y_{t+s}(\omega) \mid \mathcal{F}_t] = Y_t(\omega)$ .

We now show the first property for sake of completeness. We use the fact that  $|\sum_i A_i| \leq \sum_i |A_i|$  where all  $A_i \in \mathbb{R}$ :

$$\mathbb{E}[|Y_t(\omega)|] = \mathbb{E}[|\sum_{i=0}^n e_i(\omega)(W_{t_{i+1} \wedge t} - W_{t_i \wedge t})|] \leq \sum_{i=0}^{n-1} \mathbb{E}[|e_i(\omega)|] \times \mathbb{E}[|W_{t_{i+1} \wedge t} - W_{t_i \wedge t}|]$$

By assumption all  $|e_i(\omega)| < \infty$ . Furthermore, [increments in Brownian motion are normally distributed](#) and so  $W_{t_{i+1} \wedge t} - W_{t_i \wedge t} \sim \mathcal{N}(0, t_{i+1} \wedge t - t_i \wedge t)$ . For any r.v.  $X \sim \mathcal{N}(\mu, \sigma^2)$  we have  $\mathbb{E}[|X|] = \sigma \sqrt{\frac{2}{\pi}} \implies \mathbb{E}[W_{t_{i+1} \wedge t} - W_{t_i \wedge t}] = \sqrt{\frac{2}{\pi}(t_{i+1} \wedge t - t_i \wedge t)}$ . So:

$$\mathbb{E}[|Y_t(\omega)|] \leq \sqrt{\frac{2}{\pi}} \cdot \sum_{i=0}^{n-1} \mathbb{E}[|e_i(\omega)|] \sqrt{(t_{i+1} \wedge t - t_i \wedge t)} < \infty$$

which concludes our demonstration of the first property.

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<sup>a</sup>This slight adjustment is required as  $t + s$  is not necessarily equal to  $T = t_n$ . In the case  $t + s > T$ , we consider  $W_{t+s}$  as  $W_T$  – essentially stopping the Brownian motion.

### Exercise 3: Prove Mixture of Independent Brownian Motions is a Brownian Motion

Let  $W_t^{(1)}$  and  $W_t^{(2)}$  be two independent Brownian motions. Use Levy's Theorem to show that:

$$W_t := \rho W_t^{(1)} + \sqrt{1 - \rho^2} W_t^{(2)}$$

is also a Brownian motion for a given constant  $\rho$ .

To use Levy's Theorem (Theorem 2) to show that  $W_t$  is a Brownian motion we must show that  $\forall T > 0$ ,  $W_t$ 's quadratic variation over  $[0, T]$  is equal to  $T$ . We fix constants  $\rho \in \mathbb{R}, T \in \mathbb{R}^+$ , and make a partition  $0 < t_0 < t_1 < \dots < t_n = T$  of our interval  $[0, T]$ . Then we can define the sum of square changes of  $W_t$  to be  $Q_n(T) := \sum_{i=1}^n (\Delta W_i)^2$  where each  $(\Delta W_i)^2$  is given by:

$$\begin{aligned} (\Delta W_i)^2 &= [W_{t_i} - W_{t_{i-1}}]^2 = [\rho(W_{t_i}^{(1)} - W_{t_{i-1}}^{(1)}) + \sqrt{1 - \rho^2}(W_{t_i}^{(2)} - W_{t_{i-1}}^{(2)})]^2 = \\ &\quad \rho^2 (\Delta W_i^{(1)})^2 + (1 - \rho^2) (\Delta W_i^{(2)})^2 + 2\rho\sqrt{1 - \rho^2} \Delta W_i^{(1)} \Delta W_i^{(2)} \end{aligned}$$

By Levy's Theorem, because  $W_t^{(1)}$  and  $W_t^{(2)}$  are Brownian Motions, their quadratic variation over interval  $[0, T]$  is equal to  $T$ . So defining  $\Delta t := \max_i(t_i - t_{i-1})$  we have that the quadratic variation of  $W_t$ ,  $\lim_{\Delta t \rightarrow 0} Q_n(T) = \lim_{\Delta t \rightarrow 0} \sum_{i=1}^n (\Delta W_i)^2$  is given by<sup>a</sup>:

$$\begin{aligned}
\lim_{\Delta t \rightarrow 0} Q_n(T) &= \rho^2 T + (1 - \rho)^2 T + 2\rho\sqrt{1 - \rho^2} \lim_{\Delta t \rightarrow 0} \sum_{i=1}^n \Delta W_i^{(1)} \Delta W_i^{(2)} \\
&= T + 2\rho\sqrt{1 - \rho^2} \lim_{\Delta t \rightarrow 0} \sum_{i=1}^n \Delta W_i^{(1)} \Delta W_i^{(2)}
\end{aligned}$$

So to show that  $\lim_{\Delta t \rightarrow 0} Q_n(T) = T$ , we WTS that  $\lim_{\Delta t \rightarrow 0} \sum_{i=1}^n \Delta W_i^{(1)} \Delta W_i^{(2)} = 0$ .

We first begin by defining r.v.  $S_n := \sum_{i=1}^n \Delta W_i^{(1)} \Delta W_i^{(2)}$ . Note that because  $W_t^{(1)}$  and  $W_t^{(2)}$  are independent Brownian motions,  $\mathbb{E}[S_n] = \sum_{i=1}^n \mathbb{E}[\Delta W_i^{(1)}] \mathbb{E}[\Delta W_i^{(2)}] = \sum_{i=1}^n 0 \cdot 0 = 0$  and so  $\text{Var}(S_n) = \mathbb{E}[S_n^2]$ . We look at this  $\mathbb{E}[S_n^2]$  below, which will be helpful in establishing  $S_n \xrightarrow{\mathbb{P}} 0$ :

$$\begin{aligned}
\mathbb{E}[S_n^2] &= \mathbb{E}\left[\left(\sum_{i=1}^n \Delta W_i^{(1)} \Delta W_i^{(2)}\right)^2\right] = \sum_{i,j \in [1,n]} \mathbb{E}[\Delta W_i^{(1)} \Delta W_i^{(2)} \Delta W_j^{(1)} \Delta W_j^{(2)}] \\
&= \sum_{i,j \in [1,n]} \mathbb{E}[\Delta W_i^{(1)} \Delta W_j^{(1)}] \mathbb{E}[\Delta W_i^{(2)} \Delta W_j^{(2)}]
\end{aligned}$$

Because increments are independent in a Brownian motion  $\forall i \neq j$  and  $\forall k \in \{1, 2\}$  we have  $\mathbb{E}[\Delta W_i^{(k)} \Delta W_j^{(k)}] = \mathbb{E}[\Delta W_i^{(k)}] \mathbb{E}[\Delta W_j^{(k)}] = 0 \cdot 0 = 0$ . So we can continue simplifying  $\mathbb{E}[S_n^2]$  more:

$$\mathbb{E}[S_n^2] = \sum_{i=1}^n \mathbb{E}[(\Delta W_i^{(1)})^2] \mathbb{E}[(\Delta W_i^{(2)})^2] = \sum_{i=1}^n (\Delta t_i)^2 \leq \max_i \Delta t_i \times \sum_{i=1}^n \Delta t_i = \max_i \Delta t_i \times T$$

But then as  $\Delta t = \max_i \Delta t_i \rightarrow 0$  we have  $\mathbb{E}[S_n^2] \rightarrow 0$ . By Chebyshev's Inequality,  $\forall \epsilon > 0$ :

$$\mathbb{P}(|S_n| > \epsilon) \leq \frac{\text{Var}(S_n)}{\epsilon^2} = \frac{\mathbb{E}[S_n^2]}{\epsilon^2} \rightarrow 0 \text{ as } n \rightarrow \infty$$

and so by definition of [convergence in probability](#)  $S_n = \sum_{i=1}^n \Delta W_i^{(1)} \Delta W_i^{(2)} \xrightarrow{\mathbb{P}} 0$ . Thus,  $\lim_{\Delta t \rightarrow 0} Q_n(T) = T \implies W_t$  is a Brownian motion by Levy's Theorem.

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<sup>a</sup>Please note that  $\Delta t \rightarrow 0 \iff n \rightarrow \infty$ .